# Female Education Inequality in Turkey: factors affecting girls' schooling

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#### Abstract:

Education remains to be an important determinant of economic and social opportunities for individuals. However, within group inequalities for educational elements are not studied broadly. This paper seeks to inspect the reasons of higher inequality among females in terms of schooling distribution. Our results suggest that occupation of the household head, size and the composition of the family, and education of the parents have a significant impact on the schooling decision for girls. The paper will contribute to the literature in a twofold manner. First, it provides empirical evidence of schooling distribution in Turkey. Second, it investigates the factors that disproportionately affect schooling decisions for girls.

Keywords: inequality, education, gender

JEL: J24, I2

#### 1. Introduction

Education is essential for any individual's life chances and both within and across countries it continues to be a dividing line among different social groups. Besides, the individual benefits of education, it has positive spillovers therefore the whole economy and society benefit from a rise in human capital achieved through schooling. Either through its effects on earnings and skills or its function in transmitting the intergenerational disparities, education continues to be a pertinent aspect for studying inequality. Despite the theoretical importance of education, there have been few empirical studies looking at the distribution of it, particularly within a country. Furthermore, there is no detailed examination of schooling dispersion within certain socio-economic groups. Although, gap in educational opportunities between different socio-economic factions have been documented widely, the analysis have not been conducted for variations within the groups.

In the literature within group inequalities are examined to a large extent for income and wages<sup>1</sup>. For instance, the Mincerian earning equations in a number of countries have shown that between group inequalities contribute to the overall variance in a relatively small fraction. Nevertheless, the 'residual inequality' accounts for the majority of the dispersion and this is verified for many OECD countries, also, the within group inequality is mostly responsible for the changes over time (Katz and Autor, 1999). It has been also claimed that the within group inequalities and between group inequalities arise due to similar factors. However, the central reason for wage inequality, namely technology, works through separate channels for within and between group dynamics (Svizzero and Tisdell, 2003). Also, there are differential returns to race and gender but these are mainly analyzed by making across group comparisons, and not necessarily paying enough attention to what might generate variations for the members of the same category.

The investigation of within group inequalities for other socio-economic variables such as education and health has been lacking to a great extent. In a recent paper, it has been discussed that the world education inequality emanates mainly from within country component (Sahn and Younger, 2007)<sup>2</sup>. The decrease in cross country educational attainment is balanced by the rise inside the countries, hence world education inequality remains at high levels. Moreover, in South Africa the earnings inequality have risen among race groups but decreased between them during the period of 1980-1993 (Moll, 2000). The gender gap in education for China has declined over time; however the schooling continued to be more unequal among females (Saccone, 2008). All these studies reveal that socio-economic groups such as gender or race are very broad and usually are divided along income, regional and family background lines. Thus, it is necessary to investigate the within group differentials.

This paper will analyze the important aspects of educational dispersion among females in Turkey. Basic inequality measures all indicate that the female education distribution is more skewed compared to the males and this contributes significantly to the overall disparity. Besides, the higher inequality figures have been persistent over

<sup>&</sup>lt;sup>1</sup> For a broader discussion see, Aghion et al., (1999), and Machin and van Reenen (1998).

<sup>&</sup>lt;sup>2</sup> These results are quite similar to measurements and decompositions conducted for health inequality in the world.

time. Although, Turkey is no different than many countries for having relatively lower levels of education among girls, it exhibits certain peculiarities by having a greater education inequality among women. For Turkey, to our knowledge, there are no studies looking at the educational dispersion within groups, hence this paper will contribute to the literature firstly by measuring the extent of schooling distribution, and factors that lead to it. Our results suggest that occupation of the household head, size and the composition of the family, and education of the parents have a significant impact on the schooling decision for girls. Also, the paper aims to examine some of the policies that could contribute to foster more equal access and accomplishment of education in Turkey.

Next section will briefly summarize the literature on dispersion of education and existing studies for Turkey. Section 3 will present the findings for education inequality and the possible explanatory variables. Why these factors matter more for female schooling decisions will also be evaluated. In section 4, the data, results and their implications will be discussed. The last section will conclude.

# 2. Educational Distribution and Its Importance

Given the importance and the recent emphasis on skill premium and the fact that valuable skills are mostly acquired through schooling, education remains to be vital for examining individual income distribution. There are pervasive labour and capital market imperfections and education can help to overcome these if the opportunities for schooling are equally dispersed, or can aggravate them if education is mostly accessible and available to certain segments of the society. Moreover, it has been asserted that education inequality has a negative effect on economic growth of a country and income growth of the poor (Birdsall and Londono, 1997). This is partly due to underutilization of potential human capital in a nation. The income growth is also dependent on educational opportunities since in any labour market, earnings are highly correlated with schooling. Over years the share of income accruing to primary school graduates have decreased while the income received by people with higher education rose up significantly. For example, in Turkey, the ratio between illiterate and university graduate earnings soared, and this was accentuated for females. A female with no schooling earned 272.21 TL

while a female with higher education received 1273.64 TL<sup>3</sup> in real terms. This points to the fact that higher education levels pay higher in terms of income. Although a causal relationship cannot be inferred, it can be said that income and education are highly correlated in Turkey.

Schooling levels and allocation can also affect the earnings inequality in a country. It has been widely believed that higher educational attainment will bring lower wage inequalities; hence one way to overcome labour market related problems is to increase the schooling opportunities for everybody. There are various studies investigating the earnings differentials across levels of schooling in the Turkish case<sup>4</sup>, however, these studies do not focus on educational distribution. A more equitable allocation of schooling can decrease the gap between the returns to educations. Therefore, making education more accessible to masses and equalizing the schooling years can contribute to the earnings inequality reduction. This study, by emphasizing the Gini and quintile distribution, can facilitate a discussion on ways to improve income for different social groups.

There are numerous reasons why education could be unequally distributed in a society. Among them, gender, region, and culture are striking, and continue to affect decisions for human capital investments. The same households sometimes display different choices over the schooling between girls and boys. Therefore, household characteristics can generate dissimilar schooling preferences for the groups. This can be also seen regionally, since formal education is generally undervalued more in rural provinces compared to urban areas. Finally, the culture can have varied effects on girls and boys' educational decisions. Thus, it becomes important to measure the distribution of education –independent of its relation to income- besides the level of it in a country. Moreover, it is useful to look at the within group inequalities and the possible reasons to identify the differential affects of the same factors across and inside the gender categories.

<sup>&</sup>lt;sup>3</sup> For more explanation see Tansel (2004). The returns to education for girls are higher for all schooling categories.

<sup>&</sup>lt;sup>4</sup> See, Duygan and Guner for more explanation. All the educational categories are found as statistically significant.

There have been extensive studies on differential education opportunities among genders. King and Hill examined the barriers to women's education in developing countries, and conclude that despite the decrease in the gender gap, girls still receive less education than boys (King and Hill, 1993). One of the main factors behind this is the institutional structure where there are biases against girls' school choice, and another one is the cultural impediments and different standards applied to genders<sup>5</sup>. There is also discrimination and segregation in the labour market, which pushes women to select certain professions and get training for them. The World Bank data show that there has been a significant improvement for the girls' enrolment in schools. For example, the secondary school enrolment rate among girls in low income countries was around 26% in 1990 and rose to 41% in 2001. The school enrolment rate for boys in the same set of countries has gone up from 42% in 1990 to 51% in 2001 (World Bank, 2005). This indicates a convergence but there is still a 10% gap between the genders.

Despite the low investment in girls' education, the returns to schooling for women are found to be quite high. In some cases, the returns to secondary schooling for girls are estimated to be higher than it is for boys (Dollar and Gatti, 1999). This is also true for Turkey, where for all schooling categories; the female rate of return is higher according to the most recent data<sup>6</sup>. The biggest discrepancy is for secondary school, and vocational training. Women benefit extensively from attaining both formal schooling and vocational training while for men the returns have remained stable over time. Besides to the private returns, there are also socially negative effects of not sending girls to schooling or underinvesting in their education. The economic growth is adversely affected by gender inequality through two channels. The first is the direct influence of having lower average years of education and lower average quality of human capital in the nation. The second impact comes indirectly from the consequences gender education inequality has on child mortality, and fertility rates (King and Mason, 2001). It has been shown that as mother's education increases child mortality and fertility rates decrease. Both of these factors are extremely significant for the overall development of the low income countries. As understood from these studies, several factors can influence boys' and girls' education

<sup>&</sup>lt;sup>5</sup> For example, teachers are prejudiced against girls for their math and science skills.

<sup>&</sup>lt;sup>6</sup> Tunaer and Gulcan (2006) point out that this might be due to the lower level of educational opportunities, hence higher marginal returns for women.

options in a dissimilar manner leading to not only disparities across genders but also among the females themselves. Therefore, analyzing the education opportunities for girls is a prominent exercise, especially for developing countries where schooling for this group offers potential contribution to economic growth and individual well-being.

# 3. Factors behind Female Education in Turkey

## 3.1 Education Inequality

Turkish educational system can be mainly divided into two broad sections; formal and non-formal education. The former is defined as the regular education of individuals in a certain age group and given in schools at the same level with programs prepared for definite purposes, which includes pre-primary education, primary education, secondary education, and higher education institutions (MEB)<sup>7</sup>. Non-formal education largely aims to assist the formal institutions and offer life long learning for children and adults. Contemporarily, the primary education is compulsory and free in public schools. The starting age for primary schooling is six but in rural areas age seven is the most common one. Also, some of the pre-primary education institutions give training until the age of seven. Until 1998, compulsory education was for five years and this was extended to eight years later on. There are discussions to increase the compulsory schooling years to twelve years, yet these are not specified as concrete plans.

In the literature, it has been argued that education Gini and Theil indices are both better measures of inequalities in schooling especially when a big proportion of the population has no education<sup>8</sup>. In many developing countries, the total years of schooling for a majority of the population tend to be close to zero, especially for girls. Hence, to estimate the degree of inequality, these indices might be more useful. The next section will outline the Gini and Theil indices for education for both genders, and then move on to discuss the particular factors that might offer an explanation for the discrepancy.

<sup>&</sup>lt;sup>7</sup> For a broader overview, see; http://www.meb.gov.tr/Stats/apk2002ing/apage29 48.htm

<sup>&</sup>lt;sup>8</sup> The standard deviation figures become biased since in many countries no schooling category has a great share of the population.

The paper calculates education Gini and Theil indices for the last four decades for both males and females. One can immediately see that both measures have seen a dramatic decline over the years for both groups. Figure represents the development over time. The data is gathered from Barro-Lee dataset, which is constructed for 140 countries across the world<sup>9</sup> using the census and survey data for the period of 1960-2000. Seven levels of schooling categories are defined in the paper, which are no-schooling, partial primary, complete primary, partial secondary, complete secondary, partial tertiary, and complete tertiary. The corresponding years of education for each category is derived from the Ministry of Education. Figure 1 summarizes the results.



Figure 1. Education Inequality across Genders

One immediate feature that is also observable in some other parts of the world is the higher education inequity among females. In 1960, the schooling dispersion for women was much more skewed scoring a Gini coefficient of 0.86. The coefficient dropped down to 0.40 in 2000, which is still higher than the inequity for the whole population, and the males, but there has been considerable convergence. The Theil index gives similar results, although the decrease has been more rapid. However, the female

<sup>&</sup>lt;sup>9</sup> 107 of them have complete information for all variables and years.

Theil index started with a staggering 1.7 in 1960 and went down to 0.33 in 2000. One of the reasons for having a much higher inequality in terms of Theil index is because of the index's sensitivity to values near zero. Since at the earlier periods, there was a big chunk of population both among females and males with no education, the index has shown higher initial values and fell more dramatically than Gini index. Moreover, the survey data used for the individual level estimations<sup>10</sup> show that the Gini index is around 0.29 for males and 0.34 for females. The Theil indices are computed as 0.14 and 0.21 for males and females respectively. Once again, the higher inequality among women is evident, and needs further exploration.

The contribution of between group inequalities to overall schooling dispersion has ranged from 12% to 5% throughout the years. Table 1 portrays the decomposition of Theil index with respect to genders. As can be seen, the within components have a prime importance in determining the whole educational inequality, up to 95%. The results, derived from the 2004 Household Income and Consumption Expenditure Survey, reiterate this finding, the within group inequality accounts for 92% of the overall education difference. This point fits into the results of existing literature about the significance of residual inequality for variation in incomes. Education distribution in Turkey is more skewed among females. The higher levels of inequality among women indicate that the gender intensifies the within and between group inequalities. One of the ways to overcome the schooling dispersion is to focus on the inequalities among females, and identifying the determinants for this pattern.

 $<sup>^{10}</sup>$  We employed 2004 Household Income and Consumption Expenditure Survey to analyze the individual schooling decisions for boys and girls.

**Table 1.** Decomposition of Educational Inequality by Genders

	Between (%)	Within (%)
1960	0.12	0.88
1965	0.10	0.90
1970	0.11	0.89
1975	0.10	0.90
1980	0.13	0.87
1985	0.13	0.87
1990	0.10	0.90
1995	0.10	0.90
2000	0.06	0.94

Nonetheless, this does not mean that there is no substantial gender gap in terms of education in Turkey. Table 2 presents the school enrolment rates for boys and girls for different ctaegories, and it can be immediately seen that the biggest disparity is for the secondary school enrolment. For tertiary education the gap is smallest, hence bulk of the policies to overcome the differences between girls and boys should target primary and secondary education. Also, the average years of education for females increased from 1.05 years in 1960 to 3.95 years in 2000. For women, the world average for years of schooling was 4.31 in 1960 and 6.13 in 2000, therefore it can be said that Turkey is performing poorly in facilitating educational opportunities for women. While the educational inequality has been decreased, she nonetheless lacks behind the world standards. Especially the literacy rate among adult females, which is recorded as 79.6% in 2006, is still low compared to many developing countries.

**Table 2.** Enrolment Rates by Gender

	Total	Girls	Boys	Gender Gap
				_
Gross Enrolment	96.5%	93%	100%	7%
(Primary)				
Gross Enrolment	66.7%	57.2%	74.3%	17.1%
(Secondary)				
Gross Enrolment	21.5%	18.7%	24.3%	5.6%
(Tertiary)				

http://www.unicef.org/turkey/gr/ge21b.html

#### 3.2 Determinants of Girls' Education Decisions

In this section, we will try to identify some of the factors that can affect girls' schooling decisions prominently. Among these, we will consider the aspects that are disproportionately important for the girls. Although, various reasons have been provided in the literature why education remains to be biased towards boys, there aren't many studies which focus on why education can be unequally distributed within the gender categories. We will argue that occupation of the household head, size and the composition of the family, and education of the parents have a differential impact on girls and boys' schooling decisions. In addition to the preference given to males over females when schooling decisions are made, several socio-economic features deprive girls with particular backgrounds from education in Turkey. These factors can account for the majority of the within and between group education inequalities.

Income is seen as a crucial factor for schooling decisions since it determines the amount of available resources for each child and the family. Especially, when families are credit constrained the current income will significantly shape the family's ability to invest in children's education (Glewwe and Jacoby, 1994). A number of studies have found a positive connection between household income and schooling of children<sup>11</sup>. Certainly, in addition to the current income the families can have other financial assets and these are also expected to affect the education investment in a similar manner. We will use the incomes accruing from land and property owned to estimate the permanent income available for education expenditures. We will argue that income per capita for the household is more important for girls' schooling in Turkey particularly since there is son preference. In this case, parents allow girls to specialize in child labor while boys are able to acquire an education. The budget constraints make it less likely for the girls to attend any schooling categories, but especially higher education since the lower levels are state subsidized. Thus, the variable employed for financial assets of the family becomes total annual income per family member.

<sup>&</sup>lt;sup>11</sup> Birdsall (1985), and Parish and Willis (1993) are some examples. In societies with a son bias, Parish and Willis (1993, p. 866) note that "... one of the best things that can happen to a male, besides being born to rich, well-educated parents, is to have an older sister".

Additionally, the occupation of the household head is argued to be explanatory for the schooling decisions. Self-employment in this regard can influence the work chances for girls and boys differently. The boys in Turkish society are viewed as the successors of the family business; hence self-employment might decrease their schooling chances. On the contrary, girls might not face a challenge for allocation of their time between family work and education. However, it has been documented that education has a higher return for men in self-employment than wage-employment. This might decrease the successor effect and make boys to attend formal schooling. We will argue that self-employment has a positive effect on girls' and negative effect on boys' education attainment, especially at the middle and high school level. Therefore, girls whose parents are self-employed are more advantageous in terms of educational attainment. The measure will be a dummy variable accounting for the parent's self-employment status.

The community structure has a large impact on the education investment and differentially forms the schooling opportunities for boys and girls. Both the quality and quantity of schools are lower in rural areas in developing countries, and furthermore the perceived value of education tends to be lower in these communities (Hyde, 1993). Cost of education can be higher in rural regions given that the schools are farther away and the ratio of school quality to expenditure is lower. The rural-urban distinction can also capture the labor market chances to a certain extent. Since, farm employment is more widespread in the rural areas; education is less of a requirement. Although, this might decrease the schooling prospects for both girls and boys, we will claim that the effect is more profound for females. The already inadequate number of available jobs in the formal markets will be allocated mostly to males; therefore, the education for girls in the rural communities. Rural-urban distinction will be captured by a dummy variable in which rural area is defined as locations with populations less than twenty thousand.

Child labor in developing countries is pervasive and children are expected to undertake a range of low-skilled tasks such as taking care of siblings, and household duties. It has been shown in many studies that these types of chores tend to reduce the educational attainment (King and Hill, 1993). We will assert that the size of the family and the age of the siblings will be negatively related to girls' schooling. In Turkey, girls are the prime contributors to household errands, hence, they are expected to stay at home

and perform these duties more often than boys. Particularly, child care servicing of the sibling will be provided by the females<sup>12</sup>. Also, the number of dependent family members might increase the amount of tasks needed to be carried<sup>13</sup>. For example, in rural India the number of children aged 6-16 has a positive impact on the time used to work and a negative effect on the time used to attend school (Cigno and Rosatti, 2002). However, it should be noted that in advanced countries, family size turned out to be insignificant under most specifications. The size of the family will definitely create a pressure on the resources and cut down the finances available for each child. Besides, the effect may be higher for girls because the extended family networks are used to the advantage of boys (Lloyd and Blanc, 1996). The size of the family and the number of siblings are derived directly from the survey.

Another factor that will be influential for the children's schooling prospects is the parental education. More educated parents are more open-minded, more willing to send their kids to schools, and perceive education more worthy (King and Hill, 1993). Moreover, the incomes are largely determined by education, and in Turkey the returns to schooling is quite high. Thus, the income and wealth will be correlated with the parents' education level. We will claim that parental education is more significant for females in traditional societies. Also, mothers with more education have more bargaining power and generate a better allocation of resources towards children's human capital (Thomas, 1994). Once again, we will argue that this increases the chances of girls' schooling more because an educated mother will be less likely to have a strong preference for the son. The variable estimating parental education will look at the father's and mother's education separately.

<sup>&</sup>lt;sup>12</sup> Anderson (1988) pointed out that school attendance for females is more sensitive to the number of

children under the age of five in the household than male attendance.

13 Although, most of the effect might come from the number of siblings, elderly care can be important in the Turkish context.

# 4. Data, Methodology, and Results

#### 4.1 Data

Individual level estimations are conducted by using the data coming from the Household Income and Consumption Expenditure Survey (HICES) in 2004. The survey included 8600 households, and 36000 individuals around the country. Since our main goal is to distinguish the variables that affect the schooling decisions for girls, we first split the sample by households with children. Additionally, the study considers the children between the ages of 6-20. Although, the typical age is 7 for entering primary school in Turkey, some parents send their kids a year earlier especially when alternative pre-schooling opportunities are non-existent. Thus, we regard 6 as the starting age. This gives a sample of 9722 individuals. We will consider the primary schooling attainment and high school attainment together, the dataset has 6069 children between the age of 6 and 15. The sample is also divided according to the gender categories; there are 4871 boys and 4851 girls living with their parents given the age profiles. In Table 3, the summary statistics are provided.

**Table 3.** Mean Years of Schooling for Boys and Girls

Age	Whole Sample	Girls	Boys
6	1.25	1.27	1.23
7	1.86	1.83	1.88
8	1.95	1.94	1.95
9	1.98	1.97	1.98
10	1.97	1.97	1.97
11	1.97	1.96	1.98
12	1.96	1.95	1.97
13	1.97	1.95	1.99
14	2.75	2.76	2.74
15	3.31	3.28	3.35
16	3.67	3.66	3.68
17	4.29	4.27	4.31
18	5.12	5.11	5.14
19	5.42	4.97	5.81
20	5.27	4.75	5.87
Total	2.89	2.86	2.92

As can be seen from the table, the mean education for girls and boys are fairly similar across age cohorts. Although the averages get more divergent as the age increases still, the gender categories seem to be minor. The main differences in terms of education occur within the gender groups. For example, the Gini index for boys above the age of 18 is around 0.18, and the Theil index is approximately 0.06. The same indices are 0.23 for Gini and 0.11 for Theil when we regard the sample for girls. Both girls and boys in the same age cohort have similar levels of education but they differ from each other when we consider the within group measures. Besides, the overall Gini and Theil indices are fairly similar across groups.

When we look at the mothers' and father's mean education years for the same age cohorts, we see lower levels implying that education has become more general in Turkey from one generation to the next. The inequality indices are also higher, for fathers, the education Gini is around 0.27 and the Theil is around 0.1. For mothers, both inequality indices are also greater, Gini is approximately 0.36 while Theil is 0.21. This suggests that the previous generation might have more between group inequalities for schooling than the current one. However, the females still have a more unequally distributed education profile than their male counterparts.

The other variables that will be considered for the estimation -age of the child, composition and size of the family, parental education, self-employment, rural-urban location, logarithmic disposable income per head, and regional dummies- do not display statistically significantly different means for the girls and boys belonging to the same age cohort. However, the mean values of several variables change notably within the groups and across age cohorts. These are composition of the family, parental education, and father's self-employment. For example, the mean number of siblings for the girls in the age group 6-14, is 1.3 while this is 1.9 for girls in the age group 15-20. Also, the mothers have a mean education of 3.6 years for the second group while they have 4.1 years on average for the first group. These show that it is important to look not only between girls and boys in terms of their schooling investments but also within these groups.

# 4.2 Econometric Specification

The following section discusses the econometric specification, which is employed to regress the education on the previously defined variables. Because the final schooling level is not known for the girls and boys who are enrolled in an education institution at the time of the survey, there can be biases in the attainment model. In order to overcome this problem we will employ a censored ordered probit model. This model provides the best remedies for non-negative restriction, the discreteness, the spikes and the right-censoring (King and Lillard, 1987)<sup>14</sup>. All of these problems are present in the case of educational attainment estimations. The model is as follows:

$$E^* = \beta_{\cdot} X + \varepsilon \tag{1}$$

where  $E^*$  is the desired level of education, which is a continuous variable dependent on the observable regressors, X, and a residual term,  $\epsilon$ . Nevertheless, in practice we do not observe the desired level of education but for individuals who have completed their educational life, we see a discrete educational level. The discrete level can be written as:

$$\begin{split} E &= 0 \text{ if } E^* \leq \mu_0 \\ &= 1 \text{ if } \mu \text{ } E^* = \ \mu_1, \\ &= 2 \text{ if } \mu_1 \!\! < \!\! E^* = \mu_2, \\ &\dots \\ &= J \text{ if } \mu_{i-1} = \ E^*. \end{split}$$

Then the probability of moving from one year of education to the next will be as follows given  $\varepsilon$  is normally distributed:

<sup>14</sup> See King and Lillard for the earliest application of this model to education attainment. There have been several other studies utilizing the method later on.

15

$$\begin{split} & \text{Prob}(E=0) = \Phi(\mu_0 - \beta_i x) \\ & \text{Prob}(E=1) = \Phi(\mu_i - \beta_i x) - \Phi(\mu_0 - \beta_i x) \\ & \text{Prob}(E=2) = \Phi(\mu_2 - \beta_i x) - \Phi(\mu_1 - \beta_i x) \\ & \cdots \\ & \text{Prob}(E=J) = 1 - \Phi(\mu_{i-1} - \beta_i x) \end{split}$$

The likelihood function for uncensored observations, L<sub>u</sub>, is thus

$$\begin{split} L_u &= \Phi(\mu_E \text{--} \beta_i x) \text{ for } E = 0 \\ L_u &= \Phi(\mu_E \text{--} \beta_i x) \text{--} \Phi(\mu_{E\text{--}1} \text{--} \beta_i x) \text{ for } E = 1...J\text{--}1 \\ L_u &= 1 \text{--} \Phi(\mu_{E\text{--}1} \text{--} \beta_i x) \text{ for } S = J \end{split}$$

When we move to the censored observations, or in other words, individuals who are still enrolled, we cannot observe the completed level of education but we know that the desired level of education, E\*, is higher than the observed years, E. Therefore,

$$E^* > \mu_{E-1}$$
 which implies that  $\epsilon > \mu_{E-1}$  -  $\beta_i x$  for  $E = 0...$  J

The likelihood of the censored observations,  $L_c$ , is the probability that the error term,  $\epsilon$ , is greater than  $\mu_{E-1}$  -  $\beta_i x$  and is computed as follows:

$$L_{c} = 1 - \Phi(\mu_{E-1} - \beta_{i}x) \tag{5}$$

Finally, we multiply all the likelihood expressions, for both the uncensored and censored observations, to obtain the likelihood for the sample:

$$L = \prod L_{u} \prod L_{c}$$
 (6)

Censored ordered probit model helps us to estimate the education decisions for the individuals more accurately. Because it takes care of the non-negative restriction, the probability spikes and the discreteness of schooling, we have unbiased estimators. Additionally, we consider the likelihood functions of enrolled students and individuals who have completed their educational cycle separately, hence allowing for rightcensoring issues.

#### 4.3 Results

Table 4 summarizes the regression results for all age cohorts along the gender lines. As can be seen from the table below, there are several variables that turn out to be statistically significant. More importantly, there are several variables that are explanatory for girls but not boys. As expected age positively affects the education years of both genders, and schooling is inversely related to age squared. This is well established in the literature and confirmed by our results. Parental education is positively and significantly related to both boys and girls' education. Mother's education has higher coefficients for the girls while this is the opposite for the boys. This finding is in line with the earlier studies. For example, it has been established that mothers' education matter more for daughters in Turkey, and this is even more prominent for high school (Tansel, 2002). This is also discussed in the context of other developing countries<sup>15</sup>.

Education is a normal good, hence higher income per head brings more years of schooling. The rise in the household income decreases the probability of no schooling and increases the probability of formal education for boys and girls. This finding is unsurprising and has been evaluated already in a number of existing papers <sup>16</sup>. Financial resources are essential to for the households to be able to invest in the children's education, especially when there are no public subsidies available or the children are credit constrained. Therefore, having a higher financial resource is positively associated with more education.

<sup>&</sup>lt;sup>15</sup> Behrman (1997) provides a survey of the effect of mother's schooling on child education. The mother's education is found to be continually more important for daughters.

<sup>&</sup>lt;sup>16</sup> Birdsall (1985), Parish and Willis (1993) are among the examples.

 Table 4. Censored Ordered Probit Estimation

	Boys		Girls	
	Coefficient	(T-Statistic)	Coefficient	(T-Statistic)
Variables				
Age	0.34	(5.44)	0.24	(3.99)
Age Squared	-0.01	(-3.60)	-0.01	(-4.19)
Mother's	0.06	(3.89)	0.12	(4.28)
Education				
Father's	0.11	(4.21)	0.10	(4.05)
Education				
Log Income	0.01	(6.56)	0.03	(7.54)
Rural	-0.07	(1.53)	-0.25	(4.81)
Family Size	-0.09	(-1.16)	-0.37	(-7.75)
Family	0.06	(0.70)	-0.21	(-3.57)
Composition				
Father Self-	-0.14	(0.81)	0.32	(1.67)
employed				
Constant	2.79	(1.00)	-1.16	(-0.33)
Regions				
Aegean	0.39	(1.12)	0.95	(1.97)
Mediterranean	0.95	(2.49)	0.93	(2.31)
Central Anatolia	-0.52	(1.01)	-0.99	(3.56)
Black Sea	-1.02	(3.11)	-1.11	(3.53)
East Anatolia	-1.16	(3.33)	-1.36	(5.72)
Southeast				
Anatolia	-1.04	(3.18)	-1.16	(4.8)
Threshold Param	neters			
m1	0.04	(8.35)	0.03	(6.33)
m2	0.11	(13.45)	0.08	(9.4)
m3	0.18	(17.37)	0.14	(12.48)
m4	0.28	(21.54)	0.25	(15.7)
m5	0.49	(29.23)	0.52	(21.69)
m6	0.58	(32.17)	0.58	(22.48)
m7	0.70	(35.54)	0.63	(23.07)
m8	0.90	(39.84)	0.82	(24.62)
m9	1.00	(41.82)	0.87	(25.47)
m10	1.50	(47.01)	1.41	(26.51)
Chi-squared	17991.21		10652.08	
Significance	0.00		0.00	
Level				
Number of	4	871	48	51
Observations				
Log-likelihood	-80	82.05	-490	2.73

Living in a rural area negatively affects the education of boys and girls, nevertheless only for the female sample, the coefficient is statistically significant. As stated before urban-rural distinction might be more important for girls since the schooling opportunities are scarcer in the latter and families are more likely to keep their daughters at home. This effect is verified in Tansel (2002), and Bruce and Aytac (2004). The rural regions decrease the likelihood of having a formal education for girls in Turkey. Our analysis use a very crude measure of rural effects, and several other indicators can be inserted. For example, looking at the quality of schools in those regions, or the distance to the nearest educational facility might be helpful. Another proxy can be the teacher-pupil ratio.

The family size and family composition variables both have negative and significant coefficients for girls' sample. First measure is computed by taking the number of people living in the same household into consideration. This decreases the probability of attaining school for girls and boys; however the effect is more profound for girls. One should note that this is independent of the financial resources since the squeeze on funds is encapsulated by the income per head variable. Hence, family size matters through child labor channel or the girls are sacrificed when there are boys in the family at schooling age. Family composition is a proxy for families with small children, to check the child care service effect. The negative and significant sign for girls indicate that daughters are expected to take care of their younger siblings.

Although, both mother's and father's self-employment are put in the regression equation, only fathers are reported. There is extremely little number of mothers who are self-employed in the sample, thus we will evaluate only father's self-employment However, the effect of father being self employed was negative but not statistically significant for boys, confirming our claim that family business is expected to be passed on to sons, therefore, they need more apprenticeship than girls. Self-employment increases the chances of a girl but the variable remains to be statistically insignificant. To our knowledge, Tansel (2002) is the only other study that utilizes this measure for Turkey, and she found a negative and significant effect for middle and high school levels for boys.

When we inspect the regional dummies, we can see that except the Aegean all of them are statistically significant for girls. The reference region is Marmara. Central Anatolia, Black Sea, East Anatolia, and South East Anatolia regions are negatively related to girls' schooling. Although, being in the Aegean region increases the odd ratios of attaining higher levels of education for girls, the coefficient is not statistically significant. However, Mediterranean region has a significant and positive association with girls' schooling. These results differ from Bakis et al. (2009), in which they found a negative effect for Aegean region for instance, but the regions are divided along the statistical regional units rather than geographical units in their research.

Table 5 provides the marginal effects on the probability of attaining five years of schooling for censored ordered probit estimations. Age, age square, parental education, income per head, rural-urban distinction, family size and family composition are considered in the estimations. The magnitudes differ a lot between the two samples while the signs remain same with the exception of family composition variable. However, as can be seen from the below table, almost all variables are more powerful for girls. Therefore, to improve the female educational attainment in comparison to boys and to achieve a better distribution among the girls, these factors need to be enhanced. Provision of child and elderly care, and better educational in the rural areas will both reduce the gender gap and make schooling less dispersed within the female group.

**Table 5.** Marginal Effects on the Probability of Attaining 5 Years of Education

	Boys	Girls
Age	0.0079	0.0116
Age Squared	-0.0002	-0.0004
<b>Mothers Education</b>	0.0013	0.0057
Father's Education	0.0025	0.0046
Log Income	0.0003	0.0014
Rural	-0.0089	-0.0146
Family Size	-0.0015	-0.0117
Family	0.0020	-0.0177
Composition		
Constant	0.0001	-0.0001

#### 5. Conclusion

The paper examined the education inequality among girls in Turkey. Although, education levels continue to be different between girls and boys, and there is a gender gap, schooling is unequally distributed within the groups as well. Educational attainment generates economic and social opportunities for any individual and hence leaving certain groups outside the educational circles will limit their chances. Despite the importance of schooling, there have not been many studies on the distribution of it, particularly within socio-economic categories. In the paper, the gender schooling dispersion is investigated in more detail.

Females in Turkey display higher inequality and we considered numerous variables to explain why girls with specific characteristics are less likely to be enrolled in schools. This is supported by the over time developments in the educational inequalities between boys and girls, and within each type. The individual level estimations are conducted to distinguish the factors that can explain the girls and boys' education. The results suggest that size, and the composition of the family and education of the parents have a significant impact on the schooling decision for girls. Although, there are other features that affect the education of boys and girls in a similar manner, size and the composition of the family, parental education, and occupation of the father work differently for boys and girls.

Family size is found to be explanatory for girls but not boys, with a negative sign. The larger the family, the less likely the females will attend schools. This is not the case for boys, which might be due to a variety of reasons. Since, the financial resource allocation is controlled by the income per head; the family size impact is working through provision of child labor. This is more expected from girls residing in the household. Family composition has similar effects, decreasing the opportunities for girls to have higher education. In this case taking care of the younger sibling is the reason and this is once more demanded from girls. Finally, the self-employment of the father is found to be important and positive for girls. This is due to the perceptions about the responsible person to carry on the family business, and boys are more likely to be seen as such.

All these differences between boys and girls in terms of educational attainment indicate that the policies to enhance human capital investment should be more targeted. Although, the same policies might narrow the gap between gender groups, attention should also be paid to programs that can reduce the dispersion within these categories. Given that females have a more skewed distribution of schooling in Turkey, policies should aim to remedy for the obstacles that restrain girls with specific backgrounds to accomplish schooling.

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